



Bayesian Learning

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To be covered today

- Bayes decision theory
- Multivariate Normal Distribution
- Discriminant functions for the normal density
- Error Bounds

Materials: To read relevant portions from

- Duda and Hart, Chapter 2
- Mitchell, Chapter 6

Bayes Rule

- Consider a two category classification into two classes c_1 and c_2 .
- Let $P(c_i)$ and $p(\mathbf{x}|c_i)$ denote the prior probabilities and the class conditional probabilities respectively.
- Bayes Rule

$$P(c_i|\mathbf{x}) = \frac{p(\mathbf{x}|c_i)P(c_i)}{p(\mathbf{x})}$$

Bayes Rule (contd.)

- Naturally if we have an example \mathbf{x} such that

$$P(c_1|\mathbf{x}) > P(c_2|\mathbf{x}),$$

we would be inclined to assign class c_1 to \mathbf{x} .

- Probability of error:

$$P(\text{error}|\mathbf{x}) = \begin{array}{l} P(c_1|\mathbf{x}) \text{ if we decide } c_2 \\ P(c_2|\mathbf{x}) \text{ if we decide } c_1 \end{array}$$

- Clearly by deciding c_1 if $P(c_1|\mathbf{x}) > P(c_2|\mathbf{x})$ and c_2 otherwise, we can minimize the probability of error.

Multicategory Case

- Suppose there are k classes c_1, c_2, \dots, c_k are present thus for each of the k classes one can calculate $P_i = P(c_i|\mathbf{x})$
- Decide c_l as the class of \mathbf{x} if

$$p_l = \max p_i$$

- Out of many ways to represent a classifier, one possible way is through **discriminant functions**.
- Thus for the k classes we can calculate k discriminant functions $g_i(\mathbf{x}), i = 1, 2, \dots, k$.
- Decide the class label c_l to \mathbf{x} if

$$g_l(\mathbf{x}) > g_i(\mathbf{x}), \forall i \neq l$$

Discriminant Functions

- Discriminant functions are not unique
- We can generally replace a discriminant function $g(\mathbf{x})$ by $f(g(\mathbf{x}))$, where $f(\cdot)$ is a monotone increasing function.
- Thus for the Bayesian minimum error rate classification we can have the following equivalent discriminant functions:

$$\begin{aligned}g_i(\mathbf{x}) &= \frac{p(\mathbf{x}|c_i)P(c_i)}{p(\mathbf{x})} \\ &= p(\mathbf{x}|c_i)P(c_i) \\ &= \ln(p(\mathbf{x}|c_i)) + \ln P(c_i)\end{aligned}$$

Discriminant Funcs. (Contd.)

- Note, in the two category case, it is conventional to have a single discriminant function as we had in the case of logistic regression.
- Instead of using two different discriminant functions $g_1(\mathbf{x})$ and $g_2(\mathbf{x})$, it is more common to define a single function

$$g(\mathbf{x}) = g_1(\mathbf{x}) - g_2(\mathbf{x})$$

- Using this discriminant function we decide class c_1 if

$$g(\mathbf{x}) > 0,$$

and decide class c_2 otherwise.

The Normal Density

- Univariate normal density

$$p(x) = \frac{1}{\sqrt{2\pi}\sigma} \exp \left[-\frac{1}{2} \left(\frac{x - \mu}{\sigma} \right)^2 \right]$$

- The expected value of x for this density is

$$\mu = \mathcal{E}[x] = \int_{-\infty}^{\infty} xp(x)dx.$$

- The expected squared deviation or variance is

$$\sigma^2 = \mathcal{E}[(x - \mu)^2] = \int_{-\infty}^{\infty} (x - \mu)^2 p(x) dx.$$

The Normal Density (contd.)

- Multivariate normal density

$$p(\mathbf{x}) = \frac{1}{(2\pi)^{d/2} |\Sigma|^{1/2}} \exp \left[-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^T \Sigma^{-1} (\mathbf{x} - \boldsymbol{\mu}) \right],$$

where,

- $\mathbf{x} \in \mathcal{R}^d$
- $\boldsymbol{\mu} \in \mathcal{R}^d$ is the *mean vector*.
- Σ is the $d \times d$ *covariance matrix*.
- The above equation is often abbreviated as

$$p(\mathbf{x}) \sim N(\boldsymbol{\mu}, \Sigma)$$

The Normal Density (contd.)

- Formally we have

$$\boldsymbol{\mu} = \mathcal{E}[\boldsymbol{x}] = \int \boldsymbol{x}p(\boldsymbol{x})d\boldsymbol{x}$$

and

$$\boldsymbol{\Sigma} = \mathcal{E}[(\boldsymbol{x}-\boldsymbol{\mu})(\boldsymbol{x}-\boldsymbol{\mu})^T] = \int (\boldsymbol{x}-\boldsymbol{\mu})(\boldsymbol{x}-\boldsymbol{\mu})^T p(\boldsymbol{x})d\boldsymbol{x}$$

- **Note:** The expected value of a matrix or vector is found by taking the expected values of its components.

The Normal Density (contd.)

- Properties of Σ
 - The covariance matrix Σ is always symmetric and positive semidefinite. For our cases we shall consider Σ to be positive definite and thus $|\Sigma| > 0$.
 - The diagonal entries σ_{ii} are the variances of the respective x_i -s and the off-diagonal elements are the co-variances of x_i and x_j .
 - If x_i and x_j are statistically independent, then $\sigma_{ij} = 0$.
 - If all the off diagonal entries of Σ are 0 then $p(\mathbf{x})$ reduces to the product of the univariate densities of the components of \mathbf{x} .

DFs for Normal Density

- We saw that minimum error rate classification can be achieved by use of discriminant functions of the form

$$g_i(\mathbf{x}) = \ln p(\mathbf{x}|c_i) + \ln P(c_i)$$

- In case $p(\mathbf{x}|c_i) \sim N(\boldsymbol{\mu}_i, \Sigma_i)$, the discriminant functions can be easily evaluated. The form of the discriminant function then becomes:

$$g_i(\mathbf{x}) = -\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu}_i)^T \Sigma_i^{-1} (\mathbf{x} - \boldsymbol{\mu}_i) - \frac{d}{2} \ln 2\pi - \frac{1}{2} \ln |\Sigma_i| + \ln P(c_i)$$

DFs for Normal Density(cont.)

- We shall investigate some special cases
 - Case 1: $\Sigma_i = \sigma^2 I$ yields linear boundary
 - Case 2: $\Sigma_i = \Sigma$ yields linear boundary
 - Case 3: Arbitrary Σ_i yields hyperquadrics

Error Bounds

- Consider a two class classification scenario.
- Suppose a classifier has partitioned the feature space into two regions R_1 and R_2 corresponding to the two classes c_1 and c_2 .
- In this scenario, the probability of error would be

$$\begin{aligned} & P(\text{error}) \\ &= P(\mathbf{x} \in R_2, c_1) + P(\mathbf{x} \in R_1, c_2) \\ &= P(\mathbf{x} \in R_2|c_1)P(c_1) + P(\mathbf{x} \in R_1|c_2)P(c_2) \\ &= \int_{R_2} p(\mathbf{x}|c_1)P(c_1)d\mathbf{x} + \int_{R_1} p(\mathbf{x}|c_2)P(c_2)d\mathbf{x} \end{aligned}$$

Error Bounds (contd.)

- The probability of error can be written as

$$\begin{aligned}P(error) &= \int P(error, \mathbf{x}) d\mathbf{x} \\ &= \int P(error|\mathbf{x}) p(\mathbf{x}) d\mathbf{x}\end{aligned}$$

- Now, $P(error|\mathbf{x}) = \min[P(c_1|\mathbf{x}), P(c_2|\mathbf{x})]$
- Also we have

$$\min[a, b] \leq a^\beta b^{1-\beta}, \quad \text{for } a, b \geq 0 \text{ and } 0 \leq \beta \leq 1$$

Error Bounds (contd.)

- Combining we have

$$P(\text{error}) \leq P^\beta(c_1)P^{1-\beta}(c_2) \int p^\beta(\mathbf{x}|c_1)p^{1-\beta}(\mathbf{x}|c_2)$$

- If the conditional probabilities are normal, then we can compute the integral analytically which gives

$$\int p^\beta(\mathbf{x}|c_1)p^{1-\beta}(\mathbf{x}|c_2)d\mathbf{x} = e^{-k(\beta)}$$

Error Bounds (contd.)

$$k(\beta) = \frac{\beta(1-\beta)}{2} (\boldsymbol{\mu}_2 - \boldsymbol{\mu}_1)^T [\beta \boldsymbol{\Sigma}_1 + (1-\beta) \boldsymbol{\Sigma}_2]^{-1} (\boldsymbol{\mu}_2 - \boldsymbol{\mu}_1) + \frac{1}{2} \ln \frac{|\beta \boldsymbol{\Sigma}_1 + (1-\beta) \boldsymbol{\Sigma}_2|}{|\boldsymbol{\Sigma}_1|^\beta |\boldsymbol{\Sigma}_2|^{(1-\beta)}}.$$

- The minimum value of $e^{-k(\beta)}$ gives the **Chernoff Bound** on the error probability.
- A less sharper bound called the **Bhattacharyya Bound** is obtained by substituting $\beta = 0.5$

Issues to be addressed in next class

- In real life problems, we generally do not have access to the probability values.
- How to estimate the probabilities from data?
- Refer Chapter 3 of Duda and Hart and Chapter 6 of Mitchell
- We shall discuss such techniques in the next class